

International Conference
OPTIMIZATION AND CONTROL
Dedicated to 70th Birthday of Prof. Boris Polyak
Moscow May 19 – 21, 2005

CONVEX APPROXIMATIONS OF CHANCE CONSTRAINTS

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Overview:

- Chance constraints – motivation and difficulties
- Safe approximations of Chance constraints
- Bernstein Approximation
- Ambiguous Chance constraints

♣ Typically, optimization problems are of the form

$$\min_x \{f_0(x, \xi) : f_i(x, \xi) \leq 0, i = 1, \dots, I\} \quad (P_\xi)$$

• $x \in \mathbf{R}^n$ – decision vector • $\xi \in \mathbf{R}^N$ – data • $f_i(x, \xi)$ - given functions.

Example: Linear Programming: $f_i(x, \xi)$ are affine in x , and ξ is the collection of coefficients of these affine functions.

♣ More often than not the data in real-life optimization problems are *uncertain* – not known exactly when the problem should be solved. There are two major ways to treat this data uncertainty:

♠ Uncertain-but-bounded data: ξ runs through a given bounded “uncertainty set” $\mathcal{U} \subset \mathbf{R}^N$.

Here it is natural to require from a candidate solution to remain feasible, whatever be a realization $\xi \in \mathcal{U}$ (Robust Optimization).

♠ Stochastic data: ξ is random with a given distribution P .

Here it is natural to replace the original constraints with their *chance* versions:

$$\text{Prob}_{\xi \sim P} \{ \xi : f_i(x, \xi) > 0 \} \leq \epsilon,$$

where $\epsilon \ll 1$ is a given reliability.

$$\min_x \{f_0(x, \xi) : f_i(x, \xi) \leq 0, i = 1, \dots, I\}, \quad \xi \sim P$$

$$\Downarrow$$

$$\min_{x, \tau} \left\{ \tau : \begin{array}{l} \text{Prob}\{f_0(x, \xi) > \tau\} \leq \epsilon \\ \text{Prob}\{f_i(x, \xi) > 0\} \leq \epsilon, i = 1, \dots, I \end{array} \right\} \quad (*)$$

♣ Difficulties with chance constraints:

♠ Aside of few very specific cases, it is difficult to check whether a chance constraint, even a “linear” in ξ one:

$$p(x) \equiv \text{Prob} \left\{ f^0(x) + \sum_{j=1}^N \xi_j f^j(x) > 0 \right\} \leq \epsilon$$

is satisfied at a given point:

- when $N \gg 1$, a “closed form” analytical expression for $p(x)$ is usually out of question;
- Monte-Carlo-based estimation of $p(x)$ is prohibitively time-consuming when ϵ is really small (like 10^{-6} or less)
- ♠ The feasible set of a chance constraint usually is nonconvex, which makes optimization over this set highly problematic.

$$p(x) \equiv \text{Prob}\{f(x, \xi) > 0\} \leq \epsilon \quad (*)$$

♣ A natural way to overcome, to some extent, severe computational difficulties with chance constraint (*) is to replace it with its *computationally tractable safe approximation* – an efficiently computable convex constraint

$$F(x) \leq 0 \quad (**)$$

which is a *sufficient condition* for (*):

$$x \text{ satisfies } (**) \Rightarrow x \text{ satisfies } (*).$$

$$p(x) \equiv \text{Prob}\{f(x, \xi) > 0\} \leq \epsilon \quad (*)$$

♣ **The simplest way to build a safe approximation of chance constraint (*) with convex in x function $f(x, \xi)$ is as follows:**

♠ **Let $\phi(s) : \mathbf{R} \rightarrow \mathbf{R}$ be convex, nondecreasing and such that**

$$\phi(s) \geq 0 \quad \forall s \quad \& \quad \phi(0) \geq 1.$$

Then

$$\alpha > 0 \Rightarrow p_\alpha(x) \equiv \mathbf{E}_\xi \{ \phi(\alpha^{-1} f(x, \xi)) \} \geq p(x).$$

\Rightarrow For every $\alpha > 0$ the constraint $p_\alpha(x) \leq \epsilon$ is a safe approximation of (*), and therefore so is the constraint

$$\exists \alpha > 0 : \underbrace{\alpha p_\alpha(x) - \alpha \epsilon}_{\Psi(x, \alpha)} \leq 0. \quad (1)$$

♠ **Due to its origin, $\Psi(x, \alpha)$ is convex in $(x, \alpha > 0)$, so that the function**

$$\Phi(x) \equiv \inf_{\alpha > 0} \Psi(x, \alpha) \equiv \inf_{\alpha > 0} [\mathbf{E}_\xi \{ \alpha \phi(\alpha^{-1} f(x, \xi)) \} - \alpha \epsilon]$$

also is convex in x . It is easily seen that the (slightly weakened, as compared to (1)) convex constraint $\Phi(x) \leq 0$ is a safe approximation of (*).

$$\Phi(x) \equiv \inf_{\alpha > 0} \Psi(x, \alpha) \equiv \inf_{\alpha > 0} [\mathbf{E}_{\xi} \{ \alpha \phi(\alpha^{-1} f(x, \xi)) \} - \alpha \epsilon] \leq 0 \quad (**)$$

↓

$$\text{Prob}\{f(x, \xi) > 0\} \leq \epsilon \quad (*)$$

$[\phi : \mathbf{R} \rightarrow \mathbf{R} : \text{convex, nonnegative, nonincreasing and } \phi(0) \geq 1]$

♠ **Function $\Phi(x)$ is convex.** therefore, if $\Phi(\cdot)$ is efficiently computable (or, which is essentially the same, the expectation $\mathbf{E}_{\xi}\{\alpha^{-1}f(x, \xi)\}$ is so), convex constraint $(**)$ is a computationally tractable safe approximation of chance constraint $(*)$.

♣ **Question:** How to choose $\phi(\cdot)$?

♣ **Answer:** As far as tightness of approximation is concerned, the optimal choice of $\phi(\cdot)$ is

$$\phi(s) = \max[0, 1 + s]. \quad (!)$$

This choice results exactly in the Rockafellar's "conditional value at risk"-based upper bound on $\text{Prob}\{f(x, \xi) > 0\}$.

However: It usually is difficult to compute the function $\Phi(\cdot)$ associated with $(!)$...

$$p(x) \equiv \text{Prob}\{f(x, \xi) > 0\} \leq \epsilon \quad (*)$$

♣ **Bernstein bound:** $\phi(s) = \exp\{s\}$. Let

- $f(x, \xi)$ in $(*)$ be affine in ξ :

$$f(x, \xi) = f^0(x) + \sum_{j=1}^N \xi_j f^j(x)$$

- $\xi_1, \xi_2, \dots, \xi_N$ be independent of each other random scalars with given distributions P_j , and
- P_j possess finite explicitly computable (logarithmic) moment-generating functions:

$$\pi_j(z) = \log \left(\mathbf{E}_{s \sim P_j} \{ \exp\{zs\} \} \right) : \mathbf{R} \rightarrow \mathbf{R}.$$

For every $\alpha > 0$, we have

$$\log p(x) \leq \log (\mathbf{E} \{ \exp \{ \alpha^{-1} f(x, \xi) \} \}) = \alpha^{-1} f^0(x) + \sum_{j=1}^N \pi_j (\alpha^{-1} f^j(x))$$

\Downarrow

$$\boxed{\exists \alpha > 0 : \alpha^{-1} f^0(x) + \sum_{j=1}^N \pi_j (\alpha^{-1} f^j(x)) \leq \log(1/\epsilon)}$$

\Downarrow

$$\text{Prob}\{f(x, \xi) > 0\} \leq \epsilon$$

\Updownarrow

$$\boxed{\exists \alpha > 0 : f^0(x) + \underbrace{\sum_{j=1}^N \alpha \pi_j (\alpha^{-1} f^j(x))}_{\Psi(x, \alpha)} - \alpha \log(1/\epsilon) \leq 0}$$

\Downarrow

$$\text{Prob}\{f(x, \xi) > 0\} \leq \epsilon$$

$$\Rightarrow \left\{ \Phi(x) \equiv \inf_{\alpha > 0} \Psi(x, \alpha) \leq 0 \right\} \Rightarrow \text{Prob}\{f(x, \xi) > 0\} \leq \epsilon$$

♠ Assuming that • all $f^j(\cdot)$ are convex, and • $\xi_j \geq 0$ a.s. for all $j \geq 1$ with non-affine $f^j(x)$, function $\Psi(x, \alpha)$ is convex and efficiently computable
 \Rightarrow The constraint $\Phi(x) \leq 0$ is a safe computationally tractable approximation of the chance constraint $\text{Prob}\{f(x, \xi) > 0\} \leq \epsilon$.

$$\text{Prob} \left\{ f^0(x) + \sum_{j=1}^N \xi_j f^j(x) > 0 \right\} \leq \epsilon \quad (*)$$

$$[\xi \sim P_1 \times \dots \times P_N]$$

♣ What to do when the distributions P_j are *not* known exactly, and all we know is that the tuple $P^N = (P_1, \dots, P_N)$ of distributions of (independent of each other, and, say, bounded a.s. by 1) random variables ξ_1, \dots, ξ_N belongs to a given *-compact convex set \mathcal{P} ?

♠ Let us associate with (*) the *ambiguous chance constraint*

$$\sup_{(P_1, \dots, P_N) \in \mathcal{P}} \text{Prob}_{\xi \sim P_1 \times \dots \times P_N} \left\{ f^0(x) + \sum_{j=1}^N \xi_j f^j(x) > 0 \right\} \leq \epsilon \quad (!)$$

♠ It can be shown that the convex constraint

$$f^0(x) + \inf_{\alpha > 0} \max_{(P_1, \dots, P_N)} \underbrace{\left[\alpha \sum_{j=1}^N \log \left(\mathbf{E}_{\xi_j \sim P_j} \{ \exp\{ \alpha^{-1} f^j(x) \xi_j \} \} \right) - \alpha \log(1/\epsilon) \right]}_{\text{convex in } (x, \alpha > 0), \text{ concave in } P_j} \leq 0$$

is a safe approximation of (!). This approximation is computationally tractable, provided that \mathcal{P} is so.

$$\sup_{(P_1, \dots, P_N) \in \mathcal{P}} \text{Prob}_{\xi \sim P_1 \times \dots \times P_N} \left\{ f^0(x) + \sum_{j=1}^N \xi_j f^j(x) > 0 \right\} \leq \epsilon \quad (!)$$

♣ **Example:** Let $\mathcal{P} = \mathcal{P}_1 \times \dots \times \mathcal{P}_N$ with convex *-compact sets \mathcal{P}_j of Borel probability distributions on $[-1, 1]$. Then a safe convex approximation of (!) is given by

$$f^0(x) + \inf_{\alpha > 0} \left[\sum_{j=1}^N \alpha \hat{\pi}_j(\alpha^{-1} f^j(x)) - \alpha \log(1/\epsilon) \right] \leq 0$$

$$\hat{\pi}_j(z) = \max_{P_j \in \mathcal{P}_j} \log \left(\mathbf{E}_{s \sim P_j} \{ \exp\{sz\} \} \right)$$

♠ In many interesting cases, the function

$$\Pi^{\mathcal{P}}(z) = \max_{P \in \mathcal{P}} \mathbf{E}_{s \sim P} \{ \exp\{sz\} \}$$

is easy to compute:

\mathcal{P}	$\Pi^{\mathcal{P}}(z)$
$\{P : \text{supp}(P) \subset [-1, 1]\}$	$\exp\{ z \}$
$\left\{ P : \begin{array}{l} \text{supp}(P) \subset [-1, 1] \\ P \text{ is symmetric w.r.t. } \mathbf{0} \end{array} \right\}$	$\cosh(t)$
$\left\{ P : \begin{array}{l} \text{supp}(P) \subset [-1, 1], P \text{ is} \\ \text{unimodal w.r.t. } \mathbf{0} \end{array} \right\}$	$\frac{\exp\{ t \} - 1}{ t }$

\mathcal{P}	$\Pi^{\mathcal{P}}(z)$
$\left\{ P : \begin{array}{l} \text{supp}(P) \subset [-1, 1], P \text{ is unimodal} \\ \text{w.r.t. } \mathbf{0} \text{ and symmetric} \end{array} \right\}$	$\frac{\sinh(t)}{t}$
$\left\{ P : \begin{array}{l} \text{supp}(P) \subset [-1, 1] \\ \text{Mean}[P] = \gamma \end{array} \right\}$	$\cosh(t) + \gamma \sinh(t)$
$\left\{ P : \begin{array}{l} \text{supp}(P) \subset [-1, 1] \\ \gamma_- \leq \text{Mean}[P] \leq \gamma_+ \end{array} \right\}$	$\cosh(t) + \max[\gamma_- \sinh(t), \gamma_+ \sinh(t)]$
$\left\{ P : \begin{array}{l} \text{supp}(P) \subset [-1, 1] \\ \text{Mean}[P] = 0 \\ \text{Var}[P] \leq \sigma^2 \end{array} \right\}$	$\frac{\exp\{- t \sigma^2\} + \sigma^2 \exp\{ t \}}{1 + \sigma^2}$
$\left\{ P : \begin{array}{l} \text{supp}(P) \subset [-1, 1], P \text{ is} \\ \text{symmetric, } \text{Var}[P] \leq \sigma^2 \end{array} \right\}$	$\sigma^2 \cosh(t) + (1 - \sigma^2)$
$\left\{ P : \begin{array}{l} \text{supp}(P) \subset [-1, 1] \\ \text{Mean}[P] = \gamma \\ \text{Var}[P] \leq \sigma^2 \end{array} \right\}$	$\begin{cases} \frac{(1-\gamma)^2 \exp\{t \frac{\gamma-\sigma^2}{1-2\gamma+\sigma^2}\} + (\sigma^2-\gamma^2) \exp\{t\}}{1-2\gamma+\sigma^2}, & t \geq 0 \\ \frac{(1+\gamma)^2 \exp\{t \frac{\gamma+\sigma^2}{1+2\gamma+\sigma^2}\} + (\sigma^2-\gamma^2) \exp\{-t\}}{1+2\gamma+\sigma^2}, & t \leq 0 \end{cases}$

♣ Essentially, the Bernstein approximation deals with affine in ξ chance constraint

$$p(z) \equiv \text{Prob}_{\xi \sim P_1 \times \dots \times P_N} \{z_0 + \xi_1 z_1 + \dots + \xi_N z_N > 0\} \leq \epsilon \quad (*)$$

in variables z and states that the feasible set $\mathcal{Z}_{\text{Bern}}$ of the convex constraint

$$z_0 + \inf_{\alpha > 0} \left[\sum_j \alpha \log \left(\mathbf{E}_{\xi_j \sim P_j} \{ \exp\{\xi_j \alpha^{-1} z_j\} \} \right) - \alpha \ln(1/\epsilon) \right] \leq 0 \quad (**)$$

is inside the feasible set \mathcal{Z} of the chance constraint (*).

♣ Question: How conservative is the approximation?

♠ When interpreting the question straightforwardly:

How large should be the ratio $\epsilon/p(z)$ in order for Bernstein approximation to be able to certify the inequality $p(z) \leq \epsilon$?

the answer is “fully pessimistic”: already in the simplest case where $N = 1$ and ξ_1 takes just two values 0, 1, for every $\epsilon \in (0, 1)$ and every $T > 1$ it may happen that $p(z) < \epsilon/T$ and the Bernstein approximation is unable to certify that $p(z) \leq \epsilon$.

However: There are alternative meaningful interpretations of “level of conservatism”.

♣ Assume ξ_j are symmetrically distributed, and let us embed the original chance constraint in a natural parametric family:

$$p_\rho(z) \equiv \text{Prob}_{\xi \sim P_1 \times \dots \times P_N} \left\{ z_0 + \rho \sum_j \xi_j z_j > 0 \right\} \leq \epsilon$$

⇓

$$\mathcal{Z}[\rho] = \{z : p_\rho(z) \leq \epsilon\}$$

• $\rho \geq 0$: “noise intensity”. Assuming $\epsilon < 1/2$, the sets $\mathcal{Z}[\rho]$ shrink as noise uncertainty ρ grows.

♠ Let

$$\mathcal{Z}_{\text{Bern}}[\rho] = \left\{ z : z_0 + \inf_{\alpha > 0} \left[\sum_j \alpha \log \left(\mathbf{E}_{\xi_j \sim P_j} \{ \exp\{\rho \xi_j \alpha^{-1} z_j\} \} \right) - \alpha \ln(1/\epsilon) \right] \leq 0 \right\}$$

be the feasible sets of the Bernstein approximation. Then

$$\mathcal{Z}_{\text{Bern}}[\rho] \subset \mathcal{Z}[\rho]$$

and perhaps

$$\mathcal{Z}[\rho^+] \subset \mathcal{Z}_{\text{Bern}}[\rho] \tag{!}$$

for appropriately chosen $\rho^+ \geq \rho$. It is natural to measure the conservatism of approximation by the quantity

$$\frac{\inf\{\rho^+ : \mathcal{Z}[\rho^+] \subset \mathcal{Z}_{\text{Bern}}[\rho]\}}{\rho}$$

$$p_\rho(z) \equiv \text{Prob}_{\xi \sim P_1 \times \dots \times P_N} \left\{ z_0 + \rho \sum_j \xi_j z_j > 0 \right\} \leq \epsilon$$

$$\Downarrow$$

$$\mathcal{Z}[\rho] = \{z : p_\rho(z) \leq \epsilon\}$$

Theorem. Let ξ_j be uniformly distributed in segments $[-\sigma_j, \sigma_j]$, and let $\epsilon \leq 0.05$.
Then

$$\mathcal{Z} \left[18\sqrt{\ln(1/\epsilon)}\rho \right] \subset \mathcal{Z}_{\text{Bern}}[\rho] \subset \mathcal{Z}[\rho].$$